

**The role of attentional networks in smoking behavior among young adults:
Specific contribution of executive control.**

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Abstract

Introduction: The exploration of cognitive impairments associated with tobacco use disorder has expanded during the last decades, centrally showing working memory and executive deficits among smokers. Despite their critical role in everyday life and in the smoking cessation process, attentional abilities have seldom been explored. Previous studies yielded discordant results, and the involvement of attentional deficits in smoking habits remains unclear.

Methods: Capitalizing on the Attention Network Test, a theory-grounded task allowing the simultaneous but distinct evaluation of three attentional networks (alerting, orienting, executive control), we explored attentional abilities in three groups of 25 college students (non-smokers, light smokers, heavy smokers), matched for demographic and psychopathological characteristics.

Results: While light smokers did not present any deficit compared to non-smokers, heavy smokers showed a specific impairment of the executive control subcomponent of attention, contrasting with preserved alerting and orienting attentional abilities. The executive control deficit was not related to current craving or to smoking duration.

Conclusions: Beyond the already explored memory and executive deficits, tobacco use disorder is associated with attentional impairments, characterized by a reduced ability to focus attentional resources on pertinent stimuli and resist to distractors interference. Given the assumed role of attentional impairments in smoking, our findings suggest that a critical step in future translational iterations is to develop neuropsychological rehabilitation programs tapping into the executive network of attention among smokers.

Keywords: tobacco use disorder; attention network test; alerting; orienting; executive conflict; nicotine

Implications

This study clarifies the presence and extent of attentional impairments in tobacco use disorder. We measured three attention networks (alerting, orienting, executive control) in light smokers, heavy smokers and matched healthy controls through a theory-grounded task (Attention Network Test). Heavy smokers (but not light ones) present a specific deficit for the executive control of attention. This deficit, uncorrelated with psychopathological comorbidities or current craving, appears directly related to smoking. Given the currently scattered literature on this topic, attentional processes deserve a thorough audit in tobacco use disorder, notably to develop specific neurocognitive rehabilitation.

Introduction

Tobacco use disorder remains a key public health issue, constituting the main cause of preventable death in Western countries (WHO, 2021). Beyond the physical consequences associated with smoking habits (e.g., Didkowska et al., 2016), their psychological correlates are also established, including disinhibition (Dinn et al., 2004), extraversion (Alvarez-Lopez et al., 2001) and low self-control (Wilson & MacLean, 2013). More recent studies exploring the neuropsychological factors related to tobacco use (Conti et al., 2019 for a review), documented impaired working memory (Durazzo et al., 2010), increased impulsivity (Wagner et al., 2012) and reduced inhibition (Flaudias et al., 2016) in smokers. While acute nicotine consumption might enhance cognition (Potter & Newhouse, 2008), its chronic intake thus has long-term negative impact on high-level cognitive functions (Waisman Campos et al., 2016).

Conversely, lower-level cognitive processes, and particularly attention, have seldom been investigated. A recent meta-analysis (Conti et al., 2019) showed that only 11 studies assessed attention among smokers. This meta-analysis reported a significant but weak reduction of attentional abilities in smokers, with a wide heterogeneity across studies: Some showed slightly reduced performance in visual (Wagner et al., 2012), selective (Yakir et al., 2007) or sustained (Chamberlain et al., 2012; Friend et al., 2005) attention. Others identified such deficits in specific groups (e.g., only among non-educated smokers, Elwan et al., 1997). Finally, most studies did not identify any attentional deficit (Ernst et al., 2001; Paelecke-Habermann et al., 2013; Paul et al., 2006; Razani et al., 2004; van der Heijden et al., 2021). Such discrepancies are related to three shortcomings. First, studies were not anchored in sound attention models

articulating subcomponents in a theory-grounded investigation. Without such conceptual framework, each previous work focused on isolated attentional subcomponents (e.g., vigilance, sustained attention, divided attention), thwarting any coherent view of the impaired and preserved subcomponents' pattern. Second, previous studies did not distinguish participants according to smoking heaviness, hampering to determine the variation of attentional deficits with tobacco use severity. Third, most studies did not control for participants' socioeconomic status or other psychopathological (e.g., depression, anxiety) and substance-related (e.g., alcohol or cannabis consumption) disorders, frequently associated with smoking and impacting attentional abilities. These limitations prevent identifying the specific attentional subcomponents impaired in tobacco use disorder, and hence the development of neuropsychological rehabilitation programs.

Our study tackled these issues by measuring attentional abilities in smoking college students and distinguishing these abilities between light and heavy smokers. We relied upon the Attention Network Test (ANT; Fan et al. 2002), allowing the integrated, simultaneous and controlled evaluation of the three attentional networks included in the Posner's dominant model of attention (Petersen and Posner 2012): (1) *alerting*, the ability to reach and maintain a high level of sensitivity towards environmental stimuli, and of readiness to react to them; (2) *orienting*, the ability to select pertinent information in the environment by engaging, disengaging, and shifting attentional resources from one stimulus to another; (3) *executive control*, the ability to solve conflicts between relevant and irrelevant stimuli through top-down attentional control, thus prioritizing the stimuli pertinent to the ongoing task. The dissociation between attentional networks proposed by the ANT is supported by neuroimaging data, that identified distinct

cerebral circuits related to each network: alerting is related to superior temporal and thalamic activations, orienting relies on superior parietal lobule, prefrontal cortex, and fusiform gyrus activations; and executive control is mostly associated with precuneus, thalamic, cingulate, and superior/inferior frontal gyri activations (Fan et al., 2005; MacLeod et al., 2010; Thienel et al., 2009; Visintin et al., 2015).

The ANT has been widely used to identify attentional deficits in psychiatric disorders (e.g., Heeren et al., 2015). Specific impairments in attentional networks have also been identified in substance use disorders, including reduced orienting abilities among people presenting cannabis dependence (Ortega-Mora et al., 2021) or impaired executive control among cocaine consumers (Woicik et al. 2009), cannabis users (Epstein & Kumra, 2014) and in severe alcohol use disorder (Maurage et al., 2014). Conversely, the ANT has only been used in tobacco research to explore the acute effects of nicotine, through studies administering nicotine to smokers or non-smokers. While earlier studies, using less controlled attentional tasks, had suggested that nicotine increased attentional abilities, particularly among non-smokers (e.g., Heishman et al., 2010), studies using the ANT showed that increased nicotine levels generated shorter reaction times (RT) during cognitive tasks (Ettinger et al., 2017) but did not specifically improve the efficiency of attentional networks (Ettinger et al., 2017; Myers et al., 2013) and might even reduce orienting abilities (Wignall & de Wit, 2011). The long-term links between tobacco use disorder and attentional abilities have thus not been explored with the ANT in a well-controlled group of smokers. Measuring such link constituted our main aim. We also wanted to determine whether attentional abilities are modulated by heaviness of smoking habits, by comparing healthy controls, light smokers and heavy smokers. Given previous results (Conti et al., 2019) showing that

tobacco use disorder is associated with impaired attention (particularly for rejecting irrelevant information) and with executive control deficits, we hypothesized that smokers, and particularly heavy smokers, will mostly exhibit reduced abilities in the executive control of attention compared to healthy controls.

Methods

Participants

We first conducted a screening phase by sending an e-mail to all the students of the University (UCLouvain, Belgium, approximately 28,000 individuals), proposing them to fill in an online questionnaire on Qualtrics software. This self-reported questionnaire assessed sociodemographic characteristics (age, gender, education level, mother tongue), psychological/physiological variables (past/present neurological or psychopathological disorder, visual abilities, medication), and tobacco/alcohol use (frequency/intensity of tobacco/alcohol consumption). We obtained 4,520 complete answers, and we then selected 1,710 students fulfilling the following inclusion criteria: native French speaker, 18-25 years old, no past/present psychological or psychiatric disorder, no past/present drug consumption (except alcohol and tobacco), absence of alcohol use disorder [measured by the Alcohol Use Disorders Identification Test (AUDIT, Saunders et al., 1993)], no current use of other tobacco-related products than cigarettes (e.g., e-cigarettes, oral tobacco), no current psychotropic medication, no major neurological problem, normal or corrected-to-normal vision. We then selected three groups presenting the targeted smoking profiles: non-smoking healthy controls (HC, who had not smoked during the last year and had never been regular smokers), light smokers (LS, who smoked between two and eight cigarettes per day), and heavy

smokers (HS, who smoked between 10 and 25 cigarettes per day). All individuals who agreed to participate in the experiment (i.e., who gave their e-mail addresses during the screening phase) and corresponded to one of the three targeted profiles (178 students) were contacted to participate in the experimental phase. We recruited all available participants until our sample size (25 participants) was obtained for each group. Seventy-five college students [65.3% women; mean age: 21.80 years old (SD=2.02)] thus performed the experiment. A power computation (performed in G*Power v3.1.9.7) indicated that a total sample size of 69 was required to detect a Group x Attentional network (three groups, three measurements) interaction in a repeated-measures ANOVA, assuming a medium ($f=0.20$) effect size with 0.90 power and $\alpha=0.05$, thus suggesting that our study is sufficiently powered. All participants provided their written informed consent and this study's protocol was approved by the local ethics committee (Psychological Science Research Institute, UCLouvain) and complied with the Helsinki Declaration of 1975, as revised in 2008.

Measures and analyses

Psychopathological and tobacco-related measures

We measured psychopathological variables before starting the experiment: (a) depressive symptoms, using the Beck Depression Inventory (BDI-II; Beck et al. 1996, French validation: Beck et al. 1998), a 21-item questionnaire, each leading to a 0-3 score (total score range: 0-63); (b) anxiety, using the State-Trait Anxiety Inventory (STAI; Spielberger et al., 1983, French validation: Bruchon-Schweitzer and Paulhan 1993), proposing 20 items related to state anxiety (STAI A) and 20 items related to trait anxiety (STAI B), each leading to a 1-4 score (total score range: 20-80); (c) impulsivity,

using the UPPS-P Impulsive Behavior Scale (Cyders et al., 2014, French validation: Billieux et al., 2012), a 20-item questionnaire, each related to an impulsivity facet (positive urgency, negative urgency, lack of premeditation, lack of perseverance, sensation seeking; score range for each item: 1-4; score range for each facet: 4-16).

In both smoking groups, we also measured: (a) the duration of smoking habits, the number of cigarettes per day and the number of pack years (i.e., the packs of cigarettes smoked per day by the participant, multiplied by the length of consumption in years); (b) the heaviness of physical addiction to nicotine, using the Fagerström test (Fagerström et al., 1989, French validation: Etter et al., 1999), a 6-item questionnaire (score range: 0-10); (c) current smoking craving, using the short version of the Questionnaire on Smoking Urges (QSU; Kozlowski et al., 1996, French validation: Dethier et al., 2014), a 12-item, each related to a craving facet [Facet 1 (7 items): relief of negative affect, Facet 2 (5 items): intention/desire to smoke; score range for each item: 1-7]. We calculated a mean score for each facet and for the total score (range: 1-7).

Stimuli and Task Description

The ANT evaluated the efficiency of three attentional networks: alerting, orienting, and executive control (Fan et al. 2002). Participants had to determine as quickly and accurately as possible the direction (left or right) of a central arrow (the target) located in the middle of a horizontal line presented either at the top or bottom of the screen (Figure 1). They responded by pressing the corresponding button (left or right) on the keyboard. Each target was preceded by either no cue, a central cue (an asterisk replacing the fixation cross), a double cue (two asterisks, one appearing above and

one below the fixation cross), or a spatial cue (an asterisk appearing above or below the fixation cross, indicating the location of the upcoming target) (Figure 1; upper part). Three possible flankers also appeared horizontally on each side of the target: arrows (two on each side of the target) pointing in the same direction as the target (congruent condition), arrows pointing in the opposite direction than the target (incongruent condition), or dashes (neutral condition) (Figure 1; middle part). Each trial comprised: (1) a central fixation cross (random duration between 400 and 1600ms); (2) a cue (100ms); (3) a central fixation cross (400ms); (4) a target and its flankers, appearing above or below the fixation cross, and remaining on the screen until the participant answered or for 1700ms if no response occurred; (5) a central fixation cross (lasting for 3500ms minus the sum of the first fixation period's duration and the RT) (Figure 1; lower part). We recorded RT (in ms) and accuracy score (correct/wrong answer) for each trial.

The ANT comprised 288 trials (three blocks of 96 trials, with a short break between blocks). The combination of four cues (no cue, center cue, double cue, spatial cue), three flankers (congruent, incongruent, neutral), two directions of the target (left, right), and two localizations (upper or lower part of the screen) generated 48 possible trials. We randomized trials presentation, each trial appearing twice per block. The task was programmed and administered via E-Prime 2 Professional® (Psychology Software Tools, Pittsburgh, PA, USA).

Procedure

This study comprised three steps. First, participants were provided with full details on the aims and procedure, and filled in the questionnaires. During this step, we also

checked that participants had not smoked during the last three hours, as requested during the recruitment. This is because nicotine has a short-term impact on cognition, notably reducing RT (Ettinger et al., 2017) and potentially increasing global attentional abilities (Heishman et al., 2010). Recent smoking could thus have increased attentional abilities in smokers, biasing our results. Second, participants performed the ANT after completing a preliminary practice session (24 randomly selected trials). The distance between the participant's eyes and the screen was 50 centimeters, and the target stimuli subtended a visual angle of 4° in the horizontal field. Finally, participants were debriefed and received compensation (10€). Each session was administrated individually in a dimly lit and quiet room.

Preparation of the Data

In line with earlier ANT studies (e.g., Fan et al. 2002; Maurage et al. 2014), we excluded data from trials with incorrect answers (3.32% of trials), RT lower than 200ms or greater than 3000ms (0.12%), and RT exceeding 2 SD below or above each participant's mean for each experimental condition (2.8%). A preliminary analysis showed no difference in RT or accuracy according to the direction and localization of the arrow. We thus merged these trials, leading to 24 trials for each of the 12 experimental conditions (four cues by three flankers). Following Fan et al. (2002), we computed indices (i.e. difference scores between two experimental conditions presenting the same working memory and motor planning requirements but differing regarding attentional resources involved) for the three attentional networks, both for RT and accuracy, and for each participant individually. This approach yielded three indices: (1) *alerting*, by subtracting the mean score for the double cue trials from the mean score for the no cue trials (i.e.

No cue – Double cue); (2) *orienting*, by subtracting the mean score for the spatial cue trials from the mean score for the central cue trials (i.e. Central cue – Spatial cue); (3) *executive control*, by subtracting the mean score for congruent flanker trials from the mean score for incongruent flanker trials (i.e. Incongruent flanker – Congruent flanker). For both alerting and orienting indices, greater subtraction scores for RT (and lower for accuracy) indicated greater efficiency. Conversely, greater subtraction scores for RT (and lower for accuracy) on executive control indicated increased difficulty (Fan et al. 2005).

Statistical Analyses

We performed statistical analyses using SPSS 25.0 ® (IBM, Inc.) and set the significance level at an alpha of .05 (bilateral). First, we performed descriptive statistics and one-way analyses of variance (ANOVAs) or independent samples t-tests to explore group differences. Second, we performed two types of ANOVAs, separately for accuracy and RT: (1) 3x4x3 ANOVAs with Group (HC, LS, HS) as between-subjects factor and Cue (no cue, central cue, double cue, spatial cue) as well as Flanker (congruent, incongruent, neutral) as within-subjects factors; (2) 3x3 ANOVAs with Group (HC, LS, HS) as between-subjects factor and Attentional Network (alerting, orienting, executive control) as within-subjects factor. Significant main effects and interactions were followed by post-hoc Bonferroni-corrected independent samples t-tests. As we focused on potential deficits among smokers, the results section will present group comparisons, and the overall effects (i.e., significant results not related to group differences) are reported in the Supplementary Materials. Pearson's correlations (performed in an exploratory way and thus uncorrected) measured the

relationships between attentional performance, demographic factors, psychopathological symptoms and tobacco-related variables in the whole sample and in each group.

Results

Demographic, psychopathological, and alcohol-related measures

As detailed in Table 1, groups did not significantly differ for age, gender, depression, anxiety and alcohol consumption. We found no group difference for lack of perseverance and sensation seeking, but smoking groups scored higher than HC for negative urgency, positive urgency and lack of premeditation. Smoking groups differed regarding the number of cigarettes smoked per day, pack years and the Fagerström test, but not on the duration of smoking habits or tobacco craving.

Outcome measures

General Analysis. We performed a 3 (groups) X 4 (cues) X 3 (flankers) ANOVA separately for RT and accuracy (Table S1 in Supplementary Materials presents the full results):

- Accuracy: We found no main Group effect [$F(2,72)=1.97$, $p=0.15$, $\eta^2_p=0.05$], nor interactions with Cue [$F(6,216)=1.07$, $p=0.38$, $\eta^2_p=0.03$], Flanker [$F(4,144)=2.39$, $p=0.05$, $\eta^2_p=0.06$] or triple interaction [$F(12,432)=0.91$, $p=0.54$, $\eta^2_p=0.02$].

- RT: We found no main Group effect [$F(2,72)=1.58$, $p=0.21$, $\eta^2_p=0.04$] nor interaction with Cue [$F(6,216)=1.41$, $p=0.21$, $\eta^2_p=0.04$] or triple interaction [$F(12,432)=1.38$, $p=0.17$, $\eta^2_p=0.04$], but a Group X Flanker interaction was found [$F(4,144)=4.56$,

$p=0.002$, $\eta^2_p=0.11$]: Groups did not differ for congruent [*HC-LS*: $t(48)=0.94$, $p=0.88$, Cohen's $d=0.27$; *HC-HS*: $t(48)=1.67$, $p=0.77$, $d=0.48$; *LS-HS*: $t(48)=0.07$, $p=1$, $d=0.02$] and neutral [*HC-LS*: $t(48)=1.09$, $p=0.66$, $d=0.31$; *HC-HS*: $t(48)=1.47$, $p=1$, $d=0.42$; *LS-HS*: $t(48)=0.26$, $p=1$, $d=0.07$] trials, but HC presented shorter RT for incongruent trials than HS [$t(48)=3.07$, $p=0.047$, $d=0.89$], other group comparisons being non-significant [*HC-LS*: $t(48)=1.77$, $p=0.21$, $d=0.51$; *LS-HS*: $t(48)=0.57$, $p=1$, $d=0.16$].

Attentional Networks Analysis. We performed a 3 (groups) X 3 (attentional networks) ANOVA separately for RT and accuracy (Figure 2):

- Accuracy: We found no main Group effect [$F(2,72)=1.53$, $p=0.22$, $\eta^2_p=0.04$] nor interaction [$F(4,144)=3.05$, $p=0.05$, $\eta^2_p=0.08$].

- RT: We found a main Group effect [$F(1,72)=5.15$, $p=0.008$, $\eta^2_p=0.13$], qualified by an interaction [$F(4,144)=3.05$, $p=0.02$, $\eta^2_p=0.08$]: We found no group differences for alerting [$F(2,74)=0.01$, $p=0.99$, $\eta^2_p=0.01$] and orienting [$F(2,74)=2.94$, $p=0.06$, $\eta^2_p=0.07$] but groups differed on executive control [$F(2,74)=4.87$, $p=0.01$, $\eta^2_p=0.12$], as HC presented better executive control (i.e. reduced index for this network) than HS [$t(48)=3.52$, $p=0.008$, $d=1.02$], the other comparisons being non-significant [*HC-LS*: $t(48)=1.99$, $p=0.18$, $d=0.57$; *LS-HS*: $t(48)=1.04$, $p=0.71$, $d=0.30$].

Correlational analyses

In the whole sample, we found significant positive correlations between (a) executive control index for accuracy and sensation seeking ($r=0.254$, $p=0.028$) as well as BDI ($r=0.232$, $p=0.046$); and (b) executive control index for RT and the duration of smoking habits ($r=0.303$, $p=0.048$). We also found significant correlations (a) *in HC*, between

executive control index for RT and STAI A ($r=0.449$, $p=0.024$) as well as STAI B ($r=0.436$, $p=0.029$); (b) *in LS*, between executive control index for accuracy and sensation seeking ($r=0.487$, $p=0.014$); (c) *in HS*, between executive control index for accuracy and the Fagerström test score ($r=0.464$, $p=0.02$). We found no other significant correlation between attentional performance and demographic, psychopathological or tobacco-related characteristics, neither in the whole sample nor within each group (all $r<0.22$, $p>0.05$).

Discussion

We examined attentional abilities among smokers through the ANT, offering an integrated exploration of the three attentional networks related to Posner's model (Petersen and Posner 2012): alerting, orienting and executive control. HS presented a deficit for the executive control network of attention, compared to matched HC (together with preserved alerting and orienting abilities), while LS presented a preserved performance for all attentional networks. Groups did not differ in accuracy, as expected since the ANT is classically related to a ceiling effect in performance (e.g., Maurage et al., 2014), replicated here (mean performance: 96.68%). HS were, however, impaired in attentional control, as demonstrated by two differences with HC regarding RT: first, the general analysis showed that HS were specifically slower for incongruent stimuli: the presence of non-relevant flankers negatively impacted their processing of the target. Second, the attentional networks analysis confirmed this result by identifying a specific deficit for the attentional control network in HS. This coherent pattern indexes executive control deficit, and specifically a difficulty to resist to distractors interference: HS have difficulties to inhibit confusing and irrelevant stimuli

(i.e., incongruent flankers) to focus their attention on significant ones (i.e., the central target). In other words, HS present a reduced ability to solve the conflict between task-relevant and task-irrelevant information. This is in line with earlier results showing impaired executive performance in smokers (Flaudias et al., 2016), particularly regarding the ability to inhibit the processing of irrelevant stimuli (Riaz et al., 2021). The LS-HS comparison also showed that this deficit appears only among people presenting strong smoking habits (i.e. between 10 and 25 cigarettes per day).

Conversely, no group differences were observed for alerting and orienting, suggesting no beneficial or negative effect of smoking on the ability to (1) globally mobilize attentional resources to increase phasic alertness and vigilance before the arrival of a target, and (2) change the focus of these attentional abilities according to task requirements to orient one's attention towards an upcoming pertinent target. This result suggests a dissociation between the short-term effects of nicotine, (i.e. increased low-level attentional abilities, particularly alerting; Heishman et al., 2010), and its long-term effects (i.e. no impact on alerting and orienting). Preserved alerting and orienting abilities also show the absence of a global attentional deficit in HS (e.g., related to reduced cognitive abilities or motivation) and thus the specificity of the executive control deficits reported. The strength of our results is further supported by the fact that attentional networks indices are based on difference scores between two experimental conditions (i.e., differential index), thus excluding the possibility that group differences are due to a global slowing down in HS. Finally, results are not the consequence of a higher difficulty of the executive control subtask, as earlier studies among populations with substance abuse showed different impairment patterns [e.g., orienting deficit in cannabis users (Ortega-Mora et al., 2021)].

Our efficient group pairing on sociodemographic and psychopathological variables supports the proposal that group differences are indeed related to tobacco use disorder. Group differences were observed on some impulsivity facets, in line with previous results showing increased impulsivity in smokers (Naudé et al., 2015). However, impulsivity is not a major contributor of our results, as (1) we found no difference on impulsivity between the two groups of smokers, despite their different ANT patterns; (2) we observed no coherent pattern of correlations between impulsivity and experimental results. In the same vein, the influence of psychopathological variables on attention appears weak, which was expected as participants only exhibited sub-clinical levels of depression and anxiety. The correlations only indicated that depression and state/trait anxiety might be related to executive control indices for accuracy and RT, respectively. These correlational analyses should nevertheless be taken cautiously as they were performed on limited samples and were not adjusted for multiple comparisons.

The grouping criteria allowed a clear dissociation between the two groups of smokers regarding the number of cigarettes per day, pack years and the physical dependence (measured through the Fagerström test), but we found no differences for smoking habits duration or current craving. The differential results between LS and HS cannot be explained by a longer duration of smoking habits or by a more intense craving among HS at testing time. This matching on craving, and the total absence of correlations between craving and experimental results, are crucial as we explicitly requested participants to avoid smoking before testing, and as it had been suggested that craving could strongly contribute to attentional impairments among smokers

(Ashare et al., 2014). This proposal is not supported by our data, showing that executive control in HS is impaired independently of current tobacco craving.

It should be acknowledged that our sample comprised young smokers with installed tobacco use disorder but quite limited duration of smoking habits (and hence relatively low pack years), low heaviness of physical nicotine addiction and low use of smoking as a coping strategy to face negative affect (i.e. Facet 1 of the QSU). The observed deficit might thus be increased in young adults presenting heavier smoking habits or in older adults with longer-lasting tobacco use disorder, potentially through a higher impact on executive control abilities and/or impairment of the other networks. Moreover, we selected students without other substance use disorder to explore the specific links between smoking and attention, but substance use comorbidities are the rule rather than the exception in youth, particularly with alcohol and cannabis (Karoly et al., 2020). As these two substances impact attention, future studies should explore the cumulative effect of tobacco and other substance abuse on attentional abilities. Finally, the cross-sectional nature of our design precludes interpretations regarding the causal relationships between tobacco use and attentional deficits: smoking may directly generate attentional impairments via neurotoxic or vascular effects, but some cognitive impairments might also be present before the initiation of smoking and facilitate its emergence (Yakir et al., 2007).

Despite these limitations, our theory-grounded evaluation of attention clarifies the contradictory results reported earlier: when offering a distinct evaluation of the three attentional networks and controlling for biasing variables, the attentional deficit observed in tobacco use disorder is exclusively related to the executive control of

attention. The various patterns of results observed earlier (Conti et al., 2019) might thus be explained by the fragmented evaluation of attentional subcomponents, using multi-determined tasks without strict control of biasing factors. The present results thus bear significant implications. At the theoretical level, they reinforce the experimental validity of the theoretical model underlying the ANT, by showing that the three networks are related to distinct and independent attentional processes. They also underline that the cognitive deficits in tobacco use disorder are not exclusively related to memory and executive functions, but also encompass attention. At the clinical level, they underline that cognitive deficits are a key feature of tobacco use disorder. Neuropsychological impairments can reduce quality of life and real-life performance, but also the ability to quit smoking (McClernon et al., 2015). Reducing attentional impairments in smokers could thus act as a therapeutic lever to facilitate smoking cessation. Neuropsychological rehabilitation programs focusing on each attentional network have been developed, and those focusing on the improvement of executive control (Serino et al., 2007) could be proposed to heavy smokers, together with recently developed rehabilitation programs using neuroscience-informed tools to reduce cognitive impairments in substance use disorders (Rezapour et al., 2016; 2020).

In conclusion, this first exploration of attentional networks in tobacco use disorder using the ANT demonstrated the presence of a differential deficit between impaired executive control and preserved alerting and orienting networks in HS, participants with lighter smoking habits presenting no attentional impairment. This result constitutes an important clarification of the previously scattered literature on attentional abilities in tobacco use disorder, and paves the way for rehabilitation programs focusing on the impacted attentional subcomponents.

Data availability statement

The dataset (SPSS version) is freely available at: https://www.uclep.be/wp-content/uploads/ANT_Smokers.sav

Declaration of interests

Alexandre Heeren receives honoraria for his editorial work from Elsevier. The other authors declare no conflict of interest.

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Table

Table 1. Sociodemographic, psychopathological and tobacco-related measures among non-smoking healthy controls (HC), light smokers (LS) and heavy smokers (HS): mean (SD).

	HC (n=25)	LS (n=25)	HS (n=25)	F(2,72) value (p-value)	T(48) value (p-value) HC-LS	T(48) value (p-value) HC- HS	T(48) value (p-value) LS- HS
<i>Sociodemographic measures</i>							
Age	21.44 (2.31)	21.72 (1.99)	22.44 (2.27)	1.00 (0.37)	/	/	/
Gender ratio (female/male)	15/10	17/8	17/8	0.47 (χ^2) (0.79)	/	/	/
<i>Psychopathological measures</i>							
Depressive symptoms (BDI-II)	6.96 (4.52)	9.96 (4.62)	9.48 (5.52)	2.69 (0.07)	/	/	/
State-anxiety (STAI-A)	35.60 (10.10)	34.36 (7.62)	36.40 (8.42)	0.34 (0.71)	/	/	/
Trait-anxiety (STAI-B)	42.56 (9.08)	44.00 (8.69)	43.60 (9.89)	0.16 (0.85)	/	/	/
Alcohol consumption (AUDIT)	7.80 (4.39)	10.16 (4.47)	9.72 (4.83)	1.89 (0.16)	/	/	/
Impulsivity (UPPS-P)							

Negative urgency	8.68 (3.12)	10.92 (3.55)	10.56 (3.04)	3.44 (0.04)	2.37 (0.02)	2.16 (0.04)	0.39 (0.70)
Positive urgency	10.48 (2.58)	11.64 (3.04)	12.72 (2.01)	4.72 (0.01)	1.45 (0.15)	3.42 (0.01)	1.48 (0.15)
Lack of premeditation	6.76 (2.22)	8.64 (2.50)	9.24 (2.98)	6.27 (0.03)	2.81 (0.07)	3.34 (0.02)	0.77 (0.44)
Lack of perseverance	7.04 (2.61)	7.84 (1.97)	6.84 (2.46)	1.26 (0.29)	/	/	/
Sensation seeking	10.52 (2.54)	11.12 (2.40)	11.36 (3.46)	0.58 (0.56)	/	/	/

Tobacco-related measures

Number of cigarettes per day	/	5.28 (1.81)	13.08 (4.42)	/	/	/	8.17 (<0.001)
Duration of smoking habits (in years)	/	4.17 (2.40)	6.00 (3.49)	/	/	/	1.99 (0.05)
Pack years	/	1.08 (0.80)	4.17 (4.34)	/	/	/	3.21 (0.003)
Fagerström test score	/	1.68 (1.18)	3.24 (2.03)	/	/	/	3.32 (0.002)
Craving (QSU): Total score	/	3.98 (0.32)	3.95 (0.37)	/	/	/	0.28 (0.78)
Craving (QSU): Relief of negative affect	/	2.86 (0.57)	2.91 (0.73)	/	/	/	0.31 (0.75)
Craving (QSU): Intention/desire to smoke	/	5.55 (0.73)	5.41 (0.66)	/	/	/	0.73 (0.47)

Figures

Figure 1. Description of the Attention Network Test, presenting the four possible cues (upper part), the six possible targets (middle part), and a trial example (lower part, i.e., congruent trial preceded by a spatial cue, the correct response being “left”). Adapted from Fan and colleagues (2002).

Figure 1.

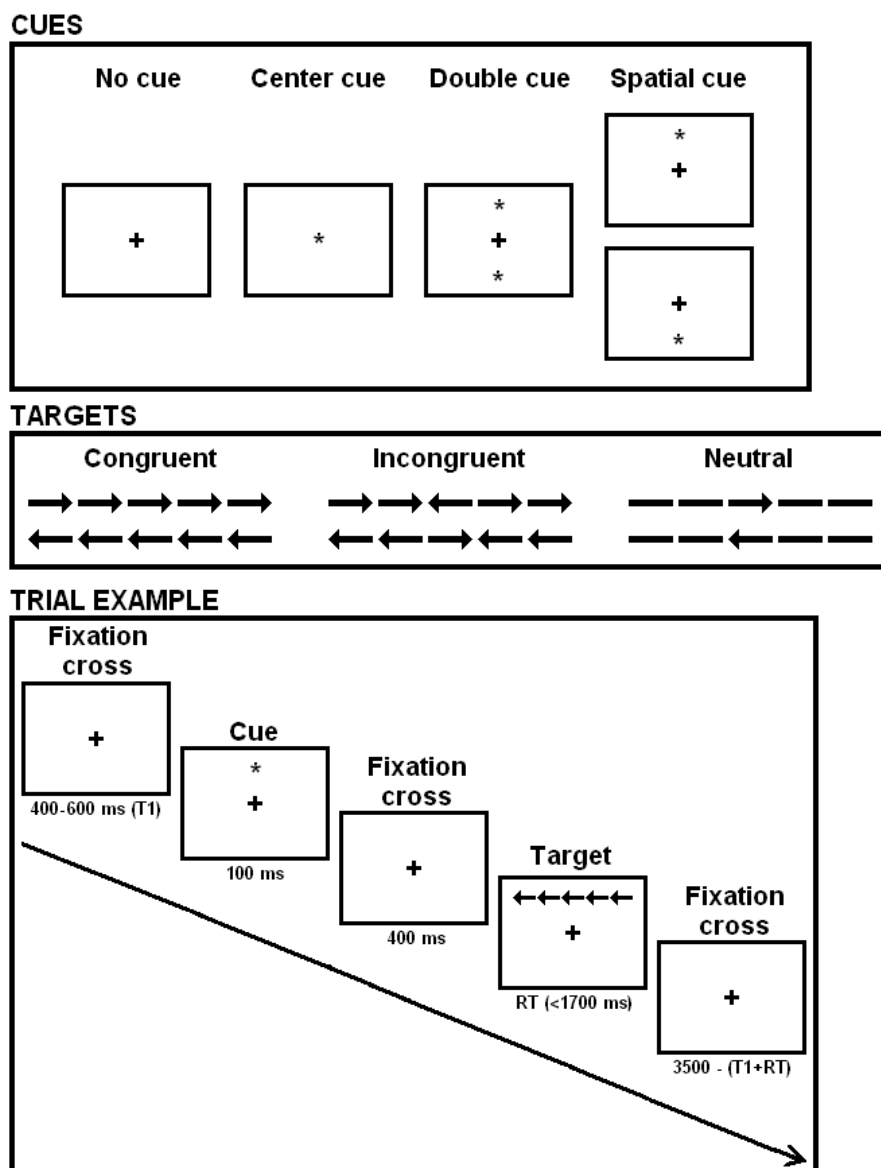


Figure 2. Differential indices (in milliseconds) for the three attentional networks (i.e., alerting, orienting, and executive control) among healthy controls (HC), light smokers (LS) and heavy smokers (HS). Columns represent the mean value for each index in each group. Error bars represent the standard errors of the mean. * $p < .01$

Figure 2.

